Recent Results in the Mathematical Modeling of Financial Bubbles

Philip Protter, Columbia University Conference in Honor of Professor Vlad BALLY Parts of this talk are based on joint work with Aditi Dandapani, others with Shihao Yang, and still other parts with Jean Jacod

October 7, 2015

How do we model financial bubbles?

- We are given a filtered complete probability space:
 (Ω, F, P, 𝔅) where 𝔅 = (𝔅_t)_{t≥0} satisfies the usual conditions, and contains at least one Brownian motion
- We let *S* denote our nonnegative stock price process, & assume interest rates are zero
- Let ${\mathbb Q}$ denote all risk neutral measures Q
- The Fundamental Price of a stock, denoted $S^* = (S_t^*)_{0 \le t \le T}$, is the conditional expectation

 $S_t^* = E_Q \{ \text{ All cash flows after time } t | \mathcal{F}_t \}$ (1)

• It is impossible really to know S_t^*

Mathematics to the Rescue

- NFLVR $\Rightarrow S_t \geq S_t^*$ a.s.
- $\beta_t = S_t S_t^* \ge 0$ is the bubble process
- Theorem[Jarrow, P², Shimbo] 2010: On a compact time interval [0, T] a stock price is undergoing bubble pricing if and only if the bubble process $\beta_t > 0$ and $\beta_t = S_t S_t^*$ is a strict local martingale under $Q \in \mathbb{Q}$
- This theorem builds on work of Lowenstein & Willard, and Cox & Hobson
- There is a lot of subsequent work by E. Bayraktar, F. Biagini, H. Föllmer, C. Kardaras, A. Nikeghbali, A. Roch, M. Schweitzer, and others

Formation of Bubbles

- We put ourselves in the framework of an incomplete market, so that there is an infinite number of risk neutral measures
- We model the risky price process using a stochastic volatility paradigm:

$$dS_t = \sigma(S_t, \nu_t) dB_t + b(S_t, \nu_t) dt$$
(2)

- Jarrow, P², and Shimbo originally proposed regime changes occurring at stopping times $T_1 \leq T_2 \leq \ldots$, where the coefficients, or especially the risk neutral measure, changes at each time T_i
- This was improved in the work of F. Biagini, H. Föllmer, and S. Nedelcu where they show that a continuous change of risk neutral measures can let the price process evolve from a martingale into a strict local martingale, thereby modeling the birth of a bubble

Indeed, to make (2) more explicit, let it be of the form

$$dS_t = \sigma(S_t, \nu_t) dB_t + \mu(S_t, \nu_t) dt$$

$$d\nu_t = f(\nu_t) dW_t + g(\nu_t) dt$$
(3)

where $d[B, W]_t = \rho dt$

- The idea of BFN is that as the risk neutral measures change, the drift of ν in (3) changes in distribution in such a way as to render S a strict local martingale
- This is quite elegant but does not directly connect bubble birth to economic reasoning
- Finally, they use the classic and renowned 1998 results of Carlos Sin, where he gives necessary and sufficient conditions for a solution of an SDE of a specific form to be a strict local martingale.

A Possible Cause for Bubbles

- What causes a stock to enter into bubble pricing (ie, speculative pricing)?
- This is the subject of many papers by economists, such as José Scheinkman & Harrison Hong
- In work with Aditi Dandapani, a PhD student at Columbia, we add information to the filtration and find that doing so can lead to bubbles
- We do this using an initial expansion technique developed by Jean Jacod in the 1980s
- More precisely, we use the work of PL Lions and M Musiela on characterizing when solutions of specific types of SDEs are strict local martingales, which they did without regard to the theory of bubbles
- The Lions-Musiela results are close to those of Sin, but are more general

 Motivated by the possibility of incorrect pricing in financial markets, Lions & Musiela studied Heston type models of the form

$$dS_t = S_t \nu_t dB_t$$

$$d\nu_t = f(\nu_t) dW_t + b(\nu_t) dt$$
(4)

where again $d[B, W]_t = \rho dt$

- We work on a time interval [0, T] (compact)
- Theorem (Lions & Musiela)

If $\limsup_{x \to \infty} \frac{\rho x f(x) + b(x)}{x} < \infty$ then S is a nonnegative martingale

If
$$\liminf_{x\to\infty} \frac{\rho x f(x) + b(x)}{\phi(x)} > 0$$
,

where $\phi(x)$ is increasing, positive, smooth, and $\int_a^{\infty} \frac{1}{\phi(x)} ds < \infty$, then S is a strict local martingale.

- We add a countable partition of events, at a time t₀ to the underlying filtration F, to get a larger filtration G
- This changes the semimartingale decompositions in (4) using (\mathbb{F}, P) and we have to remove an extra drift
- We then do a Girsanov transformation to calculate the new risk neutral measures by removing the extra drift, and we choose one we call Q, and show that under the right hypotheses (eg, ρ > 0, and with the correct assumptions on f and b in (16)) we get that S changes from a martingale under (F, P) to a strict local martingale under (G, Q)

- My thesis student at Columbia Aditi Dandapani has proved the following
- **Theorem:** Let B and W be two Brownian motions, and let S and ν satisfy

$$dS_t = S_t \nu_t dB_t$$
(5)
$$d\nu_t = f(\nu_t) dW_t + b(\nu_t) dt$$

where $d[B, W]_t = \rho dt$. Suppose that f and b are such that

$$\limsup_{x\to\infty}\frac{\rho xf(x)+b(x)}{x}<\infty, \text{ and}$$

$$\liminf_{x \to \infty} \frac{\left(\rho x f(x) + b(x) + \varepsilon f^2(x) + \varepsilon(\rho + 1) f(x)\right)}{\phi(x)} > 0$$

Then S in (5) is a (P, \mathbb{F}) martingale, and a (Q, \mathbb{G}) strict local martingale.

- An example of a choice of coefficients that works is f(x) = xand $b(x) = x - \rho x^2$
- A key tool used in the proofs of Carlos Sin, Lions-Musiela, and also Aditi, is Feller's test for explosions in the equation for ν
- We also use a relatively new concept of **locally having no** arbitrage
- This result can be extended to slightly more general frameworks, with more general coefficients

How Life Gets Messy when Data is Involved

- We want to be able to detect, from data, when pricing is in a bubble
- To begin, we choose a quite specific model
- We assume our stock price S solves an SDE of the form

$$dS_t = \sigma(S_t)dB_t + b(S_t, \nu_t)dt; \quad S_0 = 1$$
(6)

- This is an incomplete market setting
- With this framework, under any risk neutral measure *Q* equivalent to *P*, we always get the same equation:

$$dS_t = \sigma(S_t) dB_t$$

and this is key

- It is not realistic perhaps, since we do not have ν in the volatility, but it might be accurate for short amounts of time
- For equation (6) we have techniques to estimate σ(x) developed by Florens-Zmirou, and Jacod (2000). We (R. Jarrow, Y. Kchia, and P²) use a similar technique
- The (non parametric) estimate for σ has two problems:
 [a] The estimate is noisy
 [b] We can only estimate x → σ(x) for x in the range of S_t, for 0 ≤ t ≤ T

• Under $Q \in \mathbb{Q}$ we have that (2) becomes

$$dS_t = \sigma(S_t) dB_t; \quad S_0 = 1 \tag{7}$$

 We can use a Theorem of Delbaen and Shirakawa (2002): Theorem[D& S, 2002]: The process S in (7) is a nonnegative strict local martingale if and only if

[a]
$$\int_0^{\varepsilon} \frac{x}{\sigma(x)^2} dx = \infty$$
, and
[b] $\int_{\varepsilon}^{\infty} \frac{x}{\sigma(x)^2} dx < \infty$

- This theorem was improved later by Kotani and Mijatovic & Urusov
- Since S* is always a martingale, β = S S* is a strict local martingale (and hence a bubble) if and only if S is a strict local martingale, which means if and only if we have (a) & (b) above

Interpolation & Extrapolation

- We need to smooth σ to get a function to check the Delbaen-Shirakawa conditions
- We use a Reproducing Kernel Hilbert Space (RKHS) technique to smooth our estimate of σ
- The RKHS technique smooths the graph of x → σ(x) in a way analogous to using least squares to fit a line to a cloud of points; this time we fit a curve
- But this is not enough: To check (a) & (b) we need to know the behavior of x → σ(x) asymptotically as x → ∞
- Now for the part that is "louche:" we extrapolate x → σ(x) to all of [0,∞) using again an RKHS technique and an optimization criterion

Data Enters

- We got tick data for a 13 year period (2000-2013) from the Wharton Data Research Service (WRDS)
- Tick data is too noisy, so we use an idea of L. Zhang, P. Myland, and Y. Aït-Sahalia (2005) to perform a subsampling that reduces the noise
- We look for when σ behaves such that the integral

$$\int_{\varepsilon}^{\infty} \frac{x}{\sigma(x)^2} dx < \infty$$

which means we have that S is a strict local martingale and therefore $\beta > 0$, and we have a bubble

- We get a lot of false readings, and instability of the test, so we smooth the results using a Hidden Markov Model technique (HMM)
- We get a large number of fleeting bubble readings, so we impose a 5% filter: The stock price must rise more than 5% to signify the birth of a bubble, and it must later fall 5% to signify the death of a bubble, given that the test reads positive for a bubble
- The imposition of the 5% filter distorts a bit the results, and they should be interpreted with that in mind
- Using this technique, we can compute the empirical distribution of the lifetimes of financial bubbles

The Results

We get a histogram of the results which is well fit by a **generalized gamma distribution**

Histogram of duration_list\$duration



Figure: Histogram of bubble lifetimes

- We use MLE estimators to discern the parameters
- Next we try goodness of fit tests, using a QQ plot and a Kolmogorov-Smirnoff test



Nelson-Aalen estimator for cumulative hazard rate





Figure: Goodness of Fit Graphs

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The generalized gamma has the density

$$f_G(t) = \frac{\lambda p(\lambda t)^{p\kappa - 1} e^{-(\lambda t)^p}}{\Gamma(\kappa)}$$
(8)

and p and κ are shape parameters.

- If $\kappa = 1$ then the Generalized Gamma reduces to the Weibull. It also extends the log normal, the exponential, and of course the gamma distribution.
- We performed an MLE estimate for the parameters, given our extensive data set. We obtained
 - (a.) The MLE estimate for p = 0.1291201
 - (b.) The MLE estimate for $\lambda = (1/3.655156)e^{-13}$
 - (c.) The MLE estimate for $\kappa = 84.74055$

Why do we get the generalized gamma distribution?

- The generalized gamma is a bit esoteric as a distribution, unless you work in survival analysis, and are interested in the distribution of lifetimes
- In 1967, JH Lienhard and PL Meyer proposed a derivation of the generalized gamma distribution from the standpoint of problems in physics
- We can mimic their derivation, adjusting it to fit the situation of financial bubbles
- We take the convention that all bubbles begin at time t = 0.
 We can achieve this by simply translating the bubble birth time to t = 0.

- We uniformly partition \mathbb{R}_+ into intervals $[t_{i-1}, t_i)$ of length Δt .
- Next we let N_i denote the number of bubbles still alive in $[t_{i-1}, t_i)$.
- Let N be the total number of bubbles in our universe. Then

 $\frac{N_i}{N}$ is the proportion of bubbles still alive at time t_{i-1} (9)

• We assume that the proportion of bubbles alive decreases geometrically with time, and we express this as

$$\sum_{i=1}^{\infty} \left(\frac{N_i}{N}\right) t_i^{\beta} = K \text{ for constants } \beta > 0, K > 0$$
 (10)

- We also assume the death rate of bubbles alive at time t_{i-1} is proportional to a power of t_i.
- This gives that the likelihood of bubble death increases geometrically with age.
- Thus if we let g_i denote the number of bubble deaths in $[t_{i-1}, t_i)$, we assume

$$g_i = A t_i^{\alpha - 1} \tag{11}$$

so that g_i is proportional to a power of t, and the proportionality constant is A.

- We next look for the most probable distribution satisfying (9),(10) and (11).
- This gets complicated, and here we present only a sketch of the ideas.
- Let W be the number of ways bubbles can die in [t_{i-1}, t_i) given that they are alive at time t_{i-1}, for all intervals [t_{i-1}, t_i) over [0,∞).
- For example, bubbles can have a dramatic death, or they can die slowly, with a whimper, and one can give descriptions in between.
- There can also be varying economic explanations for why bubbles die, such as disagreements among different agents as to the state of current conditions; see for example the classic paper of J. Scheinkman and W. Xiong
- We obtain the following:

$$W = N! \prod_{i=1}^{\infty} \frac{g_i^{N_i}}{N_i!}$$
 (12)
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- We let \tilde{N}_i denote the values of N_i that maximize W.
- One can then show

$$\frac{\tilde{N}_{i}}{N} = \frac{\Delta t \left[\beta \left(\frac{\beta K}{\alpha}\right)^{-\alpha/\beta}\right]}{\Gamma \left(\frac{\alpha}{\beta}\right)} t_{i}^{\alpha-1} \exp\left(-\frac{\alpha}{\beta} \frac{t_{i}^{\beta}}{K}\right)$$
(13)

• The idea for showing (13) is to maximize log(W) and to use that the maximum occurs when

$$d\log(W) = \sum_{i=1}^{\infty} [\log(At_i^{lpha-1} - \log(N_i)]dN_i = 0$$

and then to use Stirling's approximation for factorial

- Our last step is to use this discrete distribution which we have found to approximate the continuous distribution
- Let τ be a stopping time. The probability that a given bubble is still alive in the interval $[t_{i-1}, t_i)$ is given by $P(t_{i-1} \leq \tau < t_i) = \tilde{N}_i/N$.
- Let the sought density f satisfy

$$rac{ ilde{N}_i}{N} = \int_{t_{i-1}}^{t_i} f(s) ds = \Delta t f(\xi)$$

by the mean value theorem, for some ξ such that $t_{i-1} \leq \xi \leq t_i$

• Next let $\Delta t
ightarrow 0$ and use (13) to get

$$f(t) = \left[\frac{\beta}{\Gamma(\alpha/\beta)} \left(\frac{\alpha}{\beta K}\right)^{\alpha/\beta}\right] t^{\alpha-1} \exp\left(-\frac{\alpha}{\beta} \frac{t^{\beta}}{K}\right), \text{ for } t \ge 0$$
(14)

where of course α, β and k are all positive (so that $f \ge 0$)

• Finally, if we make the change of variable $a = (\beta K/\alpha)^{1/\beta}$ we obtain

$$f(t) = \left(\frac{\beta}{a^{\alpha}\Gamma(\alpha/\beta)}\right)t^{a-1}\exp(-(t/a)^{\beta})$$
(15)

which is a more customary expression for the density of the generalized gamma density, and the one originally proposed by E. W. Stacy, who first proposed it in 1962

What about stochastic volatility?

Could this work if we had an equation such as

$$dS_t = \sigma(S_t, \nu_t) dB_t \tag{16}$$

under a risk neutral measure $Q \in \mathbb{Q}$?

- That is, do we have a test such as the one of Delbaen-Shirakawa to tell whether or not S is a martingale or a strict local martingale under Q ∈ Q?
- The short answer is: No.
- However we can still prove some things; what follows is work with Jean Jacod

• First we consider a solution of

$$dX_t = \sigma(X_t)dB_t; \quad X_0 = 1, \tag{17}$$

that satisfies the Delbaen-Shirakawa conditions so that X is a strict local martingale

- X in (16) is of course also a strong Markov process
- **Theorem:** X in (16) is such that $t \mapsto E(X_t)$ is strictly decreasing
- Definition: We let S denote the class of functions
 s: ℝ² → ℝ₊ such that s(x, v) = 0 if x ≤ 0 and that there exists a unique strong solution to (16) (then the solution is necessarily nonnegative, with 0 an absorbing point)

- **Definition:** For $z \ge 0$, we denote by Σ_z the class of all Borel functions σ , vanishing on $(-\infty, 0]$, positive on $(0, \infty)$, satisfying the Delbaen-Shirakawa conditions, and such that for any t > 0 and x, y > z we have $p_t^z(x, y) > 0$ for a suitable version of p_t^z .
- Proposition: Assuming s ∈ S, the solution S of (16) is a strict local martingale in the following two situations, where τ is a stopping time and τ' is an F_τ-measurable variable such that P(τ < τ', S_τ > 0) > 0:

(i) The process ν is constant (in time) on the interval $[\tau, \tau')$, and for each ν the function $x \mapsto s(x, \nu)$ belongs to Σ_0 . (ii) The process ν takes its values in some set Γ on the interval $[\tau, \tau')$, and $s(x, \nu) = \sigma(x)$ when $y \in \Gamma$, where $\sigma \in \Sigma_0$.

- The main drawback of the previous result is the fact that the time τ' is \mathcal{F}_{τ} -measurable
- We now relax this assumption, and consider a situation resembling (ii) above with Γ = (α, ∞). That is, we still assume s ∈ S, and also

$$x > \alpha, v \in \mathbb{R} \Rightarrow s(x, v) = \sigma(x).$$
 (18)

 Theorem: Assume (18) with a function σ in Σ_α ∩ Σ₀ and that the solution S of

$$dS_t = \sigma(S_t, \nu_t) dB_t; \quad S_0 = 1$$

satisfies $P(\sup_t \nu_t > \alpha) > 0$. Then S is a strict local martingale.

The End

Thank You for Your Attention